Enhancing R Control Chart Performance in Monitoring Process Dispersion using Scaled Weighted Variance Method for Skewed Populations

 ¹Abdu Atta, ²Majed Shoraim, ³Sharipah Yahya, ⁴Ali Abuzaid and ¹Esam Mahdi
 ¹Department of Mathematics, Statistics and Physics, School of Arts and Sciences, Qatar University, Doha, Qatar
 ²Department of Economics and Political Sciences, College of Commerce and Economics, Hodeidah University, Hodeidah, Yemen
 ³Department of Mathematics and Statistics, School of Quantitative Sciences, Universiti Utara Malaysia, Sintok, Malaysia
 ⁴Department of Mathematics, Faculty of Science, Al-Azhar University, Gaza, Palestine

Abstract: This study improves the performance of R control chart for monitoring process dispersion of skewed populations using scaled weighted variance method. This control chart, called Scaled Weighted Variance R control chart (SWV-R) hereafter, the SWV-R control chart compared with Skewness Correction R chart (SC-R) and Weighted Variance R chart (WV-R) in terms of false alarm. In terms of probability of detection rates the proposed SWV-R control chart reduces to the Shewhart R control chart when the underlying distribution is symmetric. An illustrative example is given to show how the proposed SWV-R control chart is constructed and works simulations study show that the proposed SWV-R control chart has the lower false alarm rates than the SC-R and WV-R control charts, when the underlying distributions are Weibull and gamma. In terms of the probability of detection rates, the proposed SWV-R control chart is closer to R control chart with the exact method than WV-R and almost the same performance as SC-R chart.

Key words: R control chart, skewed population, scaled weighted variance method, SWV-R, WV-R, SC-R

INTRODUCTION

The control charts for variables data such as the \overline{x} , EWMA, CUSUM, S and R control charts all depend on the assumption that the distribution of a quality characteristic is normal or approximately normal. However, in many situations, the normality assumption is usually violated. When the underlying distribution is non-normal, three approaches are presently employed to deal with this problem. The first approach is to increase the sample size until the sample mean is approximately normally distributed. The second approach is to transform the original data, so that, the transformed data have an approximate normal distribution. The last approach is to use heuristic methods to design control charts such as the \overline{X} and R charts based on the Weighted Variance (WV) method proposed by Bai and Choi (1995), \overline{X} control chart using Scaled Weighted Variance (SWV- \overline{x}) chart proposed by Castagliola (2000), the \overline{x} , EWMA and CUSUM charts based on the Weighted Standard Deviation (WSD) method suggested by Chang and Bai (2001), the \overline{X} and R charts based on the Skewness Correction (SC) method presented by Chan and Cui

(2003), a multivariate synthetic control chart for monitoring the process mean vector of skewed populations using weighted standard deviations suggested by Khoo et al. (2009b), a multivariate EWMA control chart using weighted variance method by Atta et al. (2014) and comparing the Median Run Length (MRL) performances of the Max-EWMA and Max-DEWMA control charts for skewed distributions by Teh et al. (2014). Other works that deal with univariate control charts for skewed distributions include that of Wu (1996), Nichols and Padgett (2005), Tsai (2007), Dou and Sa (2002), Chen (2004) and Yourstone and Zimmer (1992). In this study, the R control chart is developed by using the Scaled Weighted Variance (SWV) method suggested by Castagliola (2000). The proposed SWV-R control chart provides asymmetric limits in accordance with the direction and degree of skewness by using different variances in computing the upper and lower limits.

MATERIALS AND METHODS

The Weighted Variance (WV) method: The WV procedure splits a skewed distribution into two parts at its mean where each part is used to create a new

Corresponding Author: Abdu Atta, Department of Mathematics, Statistics and Physics, School of Arts and Sciences, Qatar University, Doha, Qatar

symmetric distribution. The two new symmetric distributions are used to set up the limits of the chart Bai and Choi (1995).

Specifically, one of the two new distributions is used to compute the upper control limit while the other is used to compute the lower control limit of the WV control chart. Since, the WV method uses a multiple of the standard deviation to establish the control limits, it requires determination of the standard deviations of the two new symmetrical distributions. Choobineh and Ballard (1987) developed a method to approximate the variance of the two distributionsas follows:

Let $\varphi(.)$ and $\Phi(.)$ denote the standard normal, N (0, 1), probability density function (pdf) and cdf, respectively. Let f(x) be pdf of quality characteristic X, from a skewed distribution; μ_x and σ_x be the mean and standard deviation of X, respectively and $P_X = P(X \le \mu_X)$. This method is based on the idea that the probability density function f(x) can be split into two new symmetrical functions, $f_{I}(x)$ and $f_{U}(x)$ having the same mean μ_X but different variances, σ_L^2 for $f_L(x)$ and σ_U^2 for $f_{\scriptscriptstyle U}(x).~f_{\scriptscriptstyle L}(x)$ and $f_{\scriptscriptstyle U}(x)$ are replaced by two normal distributions $\varphi(x, \mu_X, \sigma_L) = \varphi[(x, \mu_X) \sigma_L^{-1})]/\sigma_L$ and $\varphi(x, \mu_X, \sigma_L) = \varphi[(x, \mu_X) \sigma_L^{-1})]/\sigma_L$ $\sigma_U = \phi[(x, \mu_X) \sigma_U^{-1}]/\sigma_U$ having the same mean μ_X and variances σ_L^2 and σ_u^2 , respectively. This differs from the standard R control chart in that the standard deviation is multiplied by two different factors. One factor is used for the Upper Control Limit (UCL) while the other is used for the Lower Control Limit (LCL). Assume that $P_X =$ $P(X \le \mu_X)$ is the probability that random variable X is less than or equal to its mean μ_x . Then the UCL factor is $\sqrt{2P_x}$

and the LCL factor is $\sqrt{2(1-P_x)}$ (for more details see Choobineh and Ballard (1987).

The WV-R control chart suggested by Bai and Choi (1995) is set up by plotting the sample ranges, R_i for i = 1, 2, ..., based on the following limits:

$$UCL_{WV-R} = \mu_R + 3\sigma_R \sqrt{2P_X}$$
(1)

and:

$$LCL_{WV-R} = \mu_R - 3\sigma_R \sqrt{2(1-P_X)}$$
 (2)

where, μ_R and σ_R are the mean and standard deviation of R, respectively. Note that when $P_X = 1/2$ the WV-R control chart reduces to the standard R control chart. If the process parameters are unknown, the limits of the WV-R are:

$$\text{UCL}_{\text{WV-R}} = \overline{R} \left[1 + 3 \frac{\dot{d_3}}{\dot{d_2}} \sqrt{2 \left(1 - \hat{P}_X \right)} \right] = \text{VU}_L \overline{R}$$
(3)

and:

$$LCL_{WV-R} = \overline{R} \left[1 - 3 \frac{d_3}{d_2} \sqrt{2 \left(1 - \hat{P}_X \right)} \right]^+ = V_L \overline{R}$$
(4)

here, d_2' and d_3' are control limits constants given in Bai and Choi (1995), \overline{R} was the average of the sample ranges estimated from r preliminary subgroups and:

$$\hat{\mathbf{P}}_{\mathbf{X}} = \frac{\sum_{i=1}^{m} \sum_{j=1}^{n} \mathbf{I}\left(\overline{\overline{\mathbf{X}}} \cdot \mathbf{X}_{ij}\right)}{m \times n}$$
(5)

where, m and n are the number of samples in the preliminary data set and the sample size, respectively and I(x) = 1 if $x^3 \ge 0$ and I(x) = 0, otherwise.

Skewness Correction (SC) method: Chan and Cui (2003) proposed the Sc and R charts using Skewness Correction (SC) Method chart based on the Cornish-Fisher expansion. The limits of the SC-R chart are given by:

$$\text{UCL}_{\text{SC-R}} = \mu_{\text{R}} + \left(3 + \frac{4\alpha_3(\text{R})}{1 + 0.2\alpha_3^2(\text{R})}\right)\sigma_{\text{R}}$$
(6)

and:

$$LCL = \left[\mu_{R} + \left(-3 + \frac{4\alpha_{3}(R)}{1 + 0.2\alpha_{3}^{2}(R)} \right) \sigma_{R} \right]^{+}$$
(7)

Here, $\alpha_3(R)$ denotes the skewness of R. When the exact values of the process parameters are unknown, the control limits of the SC-R chart are Chan and Cui (2003):

$$\text{UCL}_{\text{SC-R}} = \left[1 + \left(3 + d_4^*\right) \frac{d_3^*}{d_2^*}\right] \overline{R}$$
(8)

and:

$$LCL_{SC-R} = \left[1 + \left(-3 + d_4^*\right) \frac{d_3^*}{d_2^*}\right]^+ \overline{R}$$
(9)

where, $d_4^* = \frac{4\hat{\alpha}_3(\mathbf{R})}{1+0.2\hat{\alpha}_3^2(\mathbf{R})}$, $\overline{\mathbf{R}}$ is the average sample ranges and values of d_2^* and d_3^* are given in Chan and Cui (2003).

Scaled Weighted Variance (SWV) method: Castagliola (2000) suggested an alternative approach, called the scaled weighted variance method to improve the performance of the weighted variance method. The functions $f_L(x)$ and $f_U(x)$ are not simply replaced by two normal probability density distributions $\varphi(x, \mu_X, \sigma_L)$ and $\varphi(x, \mu_X, \sigma_U)$ but are replaced by two "bell-shaped" functions $\varphi(x, \mu_X, \sigma_L, 2P_X)$ and $\varphi(x, \mu_X, \sigma_U^2)$ for second central moments and $2P_X$ and $2(1-P_X)$ for areas. Castagliola (2000) defined the function $\varphi(x, \mu_X, t, k)$ as $\phi(x, \mu_X, t, k) = \frac{k^{3/2}}{t} \varphi\left(\frac{(x-\mu_X)\sqrt{k}}{t}\right)$.

This function has the following required properties (Castagliola (2000)) for more details about the derivations:

$$\int_{-\infty}^{+\infty} \phi(\mathbf{x}, \mu_{\mathbf{X}}, t, \mathbf{k}) d\mathbf{x} = \mathbf{k}$$
(10)

$$\int_{-\infty}^{+\infty} (x - \mu_X)^2 \phi(x, \mu_X, t, k) dx = t^2$$
(11)

Using $\varphi(x, \mu_x, t, k)$ instead of the probability density function $\varphi(x, \mu_x, t)$ gives new limits for the weighted variance S control chart proposed by Khoo *et al.* (2009a). Here, the limits of the proposed Scaled Weighted Variance S control chart (SWV-S) are:

$$LCL_{SWV-R} = \mu_{R} + \Phi^{-1} \left(1 - \frac{\alpha}{4(1 - P_{X})} \right) \sqrt{\frac{P_{X}}{(1 - P_{X})}} \sigma_{R}$$
(12)

and:

$$LCL_{SWV-R} = \mu_{R} + \Phi^{-1} \left(1 - \frac{\alpha}{4P_{X}} \right) \sqrt{\frac{\left(1 - P_{X} \right)}{P_{X}}} \sigma_{R}$$
(13)

where, μ_R and σ_R are the mean and standard deviation of the R, respectively and α is Type I error rate (False alarm). Note also that when $P_X = 1/2$, the SWV-R control chart reduces to the Shewhart R control chart. If the process parameters are unknown, the control limits of the proposed SWV-R control chart are computed as follows:

$$\text{UCL}_{\text{SWV-R}} = \overline{R} \left[1 + \Phi^{-1} \left(1 - \frac{\alpha}{4 \left(1 - \hat{P}_{\text{X}} \right)} \right) \frac{d_3}{d_2} \sqrt{\frac{\hat{P}_{\text{X}}}{\left(1 - \hat{P}_{\text{X}} \right)}} \right]$$
(14)

and:

$$LCL_{SWV-R} = \overline{R} \left[1 - \Phi^{-1} \left(1 - \frac{\alpha}{4\hat{P}_{X}} \right) \frac{d_{3}'}{d_{2}'} \sqrt{\frac{\left(1 - \hat{P}_{X} \right)}{\hat{P}_{X}}} \right]^{+}$$
(15)

Here, the constants d_2' and d_3' can be computed analytically or by numerical integration. Their parameters determined for each computed via. simulation using SAS software when the underlying distributions are skewed

and $\overline{R} = \frac{\sum_{i=1}^{r} R_i}{r}$ is the average of the sample ranges

estimated from r preliminary subgroups.

Exact method of R chart for the exponential distribution: When the distribution is known, the control limits for a given Type I risk can sometimes be derived analytically. Here, we consider the case when the distribution is exponential with known parameter $\lambda = \sigma_0$ (i.e., Weibull with $\beta = 1$) (Chan and Cui (2003)) for more details. The density and distribution functions of R are:

$$f(r) = \frac{n-1}{\sigma_0} e^{-r/\sigma_0} \left[1 - e^{r/\sigma_0} \right]^{n-2}$$
(16)

and:

$$\mathbf{F}(\mathbf{r}) = \left[1 - e^{-\mathbf{r}/\sigma_0}\right]^{\mathbf{n}-1}$$
(17)

Respectively the mean and standard deviation of R can be derived analytically (Chan and Cui (2003)) and the value of skewness α_3 can be obtained by numerical integration. The control limits UCL_R and LCL_R of the R charts by the exact methods can be then obtained. The control limits are the same as those in Section 2 with known parameters. But the observations in the subgroups are from the exponential distribution with the following density function:

$$\frac{1}{(\lambda\sigma_{0})}e^{\frac{-\left[(x-(1-\lambda)\sigma_{0})\right]}{(\lambda\sigma_{0})}}, x \ge (1-\lambda)\sigma_{0}$$
(18)

An illustrate example: The data in Table 1 are generated from a gamma distribution with shape parameter, $\beta = 0.98$

Table 1: An example of illustration using simulated data from a skewed population (gamma distribution)

Sample ,i	X1	X ₂	X ₃	X_4	X ₅	\mathbf{R}_{i}	$\overline{\mathbf{X}}_{\mathrm{i}}$				
1	36.34818	22.17544	6.859294	38.06268	56.00602	49.14672	31.89032				
2	3.803384	3.073572	108.6083	57.38147	39.00405	105.5347	42.37415				
3	53.68370	50.82171	139.9070	68.93220	1.715175	138.1918	63.01195				
4	69.88040	13.33184	11.46414	0.532185	2.836212	69.34822	19.60896				
5	9.754806	29.44874	31.30970	8.411457	13.05252	22.89824	18.39545				
6	13.69733	0.156438	22.04482	10.48240	63.68418	63.52774	22.01303				
7	0.731333	15.71949	47.42605	4.871906	37.75516	46.69472	21.30079				
8	10.46658	7.359650	0.744235	0.338600	192.5089	192.1703	42.28359				
9	27.92345	2.565996	24.89902	11.83283	1.356413	26.56704	13.71554				
10	48.25964	4.733874	52.36500	19.94588	3.989499	48.37550	25.85878				
11	45.54962	73.22100	19.48764	18.21503	119.6577	101.4427	55.22620				
12	5.089729	59.90313	14.41314	9.841913	63.70672	58.61699	30.59093				
13	10.48647	78.25828	37.88342	30.91023	149.9729	139.4865	61.50227				
14	9.184076	4.387413	5.775237	114.6605	24.35253	110.2731	31.67196				

Table 1: Co	Observed valu	ies					
Sample i	 X	X	 X	x	 X	R	$\overline{\mathbf{X}}_{i}$
15	10 29390	23 09588	5 604623	10 95006	51 29480	45 69018	20 24785
16	58 14751	15,92276	42.08211	1.022873	47 96474	57.12463	33 02800
17	77.80144	57.39865	37.56566	37.30812	119.8208	82.51272	65.97894
18	18.38461	60.30539	38.73632	55.00603	50.30109	41.92078	44.54669
19	53.69745	1.597253	33.21739	20.07705	8.381358	52.10020	23.39410
20	28.38292	18.02485	24.47566	15.74064	52.77296	37.03232	27.87941
21	104.7283	6.583657	15.66652	3.788275	8.947521	100.9400	27.94285
22	21.29227	36.70789	74.14813	14.69886	33.40366	59.44927	36.05016
23	11.31113	18.36397	13.27054	49.26539	0.007235	49.25815	18.44365
24	6.521320	7.717710	2.481529	15.99499	66.52404	64.04251	19.84792
25	30.31566	1.008256	3.476084	66.72805	42.92361	65.71980	28.89033
26	24.93475	0.570747	3.297847	18.43215	23.09530	24.36400	14.06616
27	53.42357	80.60140	31.23386	1.746260	15.61345	78.85514	36.52371
28	19.82378	88.34585	9.922032	25.34298	19.09469	78.42382	32.50587
29	12.80654	18.63652	7.658047	7.148106	35.75994	28.61183	16.40183
30	24.91324	2.488491	16.33146	13.29951	3.479776	22.42475	12.10250
						$\overline{\mathbf{p}}$ - 68 69	$\overline{\mathbf{v}}$ - 31 24313

J. Eng. Applied Sci., 15 (6): 1508-1514, 2020



Fig. 1: R type control charts for the SWV, WV and SC methods using simulated data from distribution

and scale parameter, $\lambda = 40.50$. The data consist of 150 skewed observations grouped into 30 subgroups of size n = 5 each. These data are supposed to correspond to an in-control process. Since, the shape parameter, β is chosen to be 0.98 then the skewness, $\propto_3 = 2$. From these data, we compute, $\hat{\sigma}_x = 25.44$, $\hat{\mu}_x = 31.24$, $d_2' = 2.21$, $d_{3}' = 1.16$, $d_{2}^{*} = 1.41$ and $\overline{R} = 68.69$. It is observed that 95 observations fall below. Thus, $\hat{P}_x = 0.63$ using Eq. 5. Consider that $\alpha = 0.0027$, the SWV-R chart's control limits computed using Eq. 14 and 15 are equal to $UCL_{SWV-R} = 205.456$ and $LCL_{SWV-R} = -16.127$. The control limits of SWV-R chart are compared with those obtained for the WV-R control limits using Eq. 3 and 4, $UCL_{WV-R} = 190.103$ and $LCL_{WV-R} = -24.356$ and the SC-R control limits which are computed using Eq. 8 and 9, UCL_{SC-R} = 227.690 and LCL_{SC-R} = 11.363. From Figure 1, we observe that all points fall within control limits of the SWV-R and SC-R control charts indicating that the process is in-control while one points are outside the WV-R chart Upper control limit, potentially signaling a false alarm. To evaluate the performance of each of these charts, a simulation study is undertaken and the findings are further discussed.

RESULTS AND DISCUSSION

Performance evaluation and discussion of the proposed SWV-R control chart: The SWV-R control chart is compared with the SC-R and WV-R control chart for skewed data. A Monte Carlo simulation is conducted using SAS 9.4 to compute the false alarm rates and Probabilities of out-of-control detections. The false alarm rate of a control chart is defined as the proportion of subgroup points plotting beyond the limits of the chart, given that the process is actually in-control detection measures the ability of a chart in responding to a shift in the process and it represents the proportion of subgroup points plotting beyond the limits of the chart when the process has shifted. All the charts considered in this study are designed based on an in-control Average Run Length

(ARL) of 370 or Type I error of 0.0027. A shift in the process standard deviation is represented by $\sigma_1 = \delta \sigma_x$ where $\delta \in \{1.1, 1.2, 1.3, 1.4, 1.5, 2.0, 2.5, 3.0, 3.5, 4.0, 4.5\}$ is the magnitude of a shift in process standard deviation. The skewed distributions considered here are Weibull and gamma because they represent a wide variety of shapes from symmetric to highly skewed. For the sake of comparison, the standard normal distribution is also considered. For convenience, a scale parameter of one is used for the Weibull and gamma distributions. Note that P_x for the Weibull and gamma distributions are:

$$P_{\rm X} = 1 - \exp\left[-\left(\Gamma\left(1 + \frac{1}{\beta}\right)\right)^{\beta}\right]$$
(19)

and:

$$P_{\rm X} = F(\gamma) \tag{20}$$

where β and γ are the shape Respectively parameters . Here, $\Gamma(.)$ is the gamma function while F(.)is the gamma distribution functions, respectively. In the case of the false alarm rates, the skewness coefficients considered are $\alpha_3 \in \{0.5, 1.0, 1.5, 2.0, 2.5, 3.0\}$ while skewness coefficient, $\alpha_3 = 2$ is considered in the case of the probability of out-of-control detection. The sample sizes, $n \in \{5, 7, 10\}$ are considered. The false alarm rate and probability of out-of-control detection are obtained based on 10000 simulation trials. The simulated results are tabulated in Table 2 and 3 for the false alarm rate in the cases of known and unknown parameters while probability of out-of-control detections are tabulated in Table 4-6, respectively where the smallest value are bolded. Table 2 and 3 show that the proposed SWV-R control chart has lower false alarm rate than the SC-R and WV-R control charts for almost all levels of

Table 2: False alarm rates of the SWV-R, SC-R and WV-R control charts for known parameters Sample size (n)

		5			7			10		
Distribution	α3	SWV-R	SC-R	WV-R	SWV-R	SC-R	WV-R	SWV-R	SC-R	WV-R
Normal	0.0	0.0064	0.0021	0.0241	0.0059	0.0025	0.0088	0.0055	0.0032	0.0047
Weibull										
2.2266	0.5	0.0032	0.0018	0.0033	0.0025	0.0024	0.0032	0.0020	0.0024	0.0024
β 1.5688	1.0	0.0045	0.0045	0.0058	0.0040	0.0057	0.0043	0.0036	0.0048	0.0040
1.2123	1.5	0.0064	0.0099	0.0082	0.0058	0.0134	0.0067	0.0051	0.0085	0.0062
0.9987	2.0	0.0082	0.0166	0.0103	0.0073	0.0246	0.0084	0.0062	0.0118	0.0073
0.8598	2.5	0.0098	0.0222	0.0114	0.0089	0.0372	0.0099	0.0076	0.0143	0.0077
0.7637	3.0	0.0108	0.0250	0.0118	0.0099	0.0436	0.0107	0.0083	0.0167	0.0083
gamma										
15.4	0.5	0.0060	0.0056	0.0063	0.0056	0.0041	0.0066	0.0052	0.0048	0.0057
3.913	1.0	0.0066	0.0088	0.0085	0.0060	0.0072	0.0071	0.0057	0.0064	0.0069
α 1.788	1.5	0.0074	0.0097	0.0099	0.0067	0.0126	0.0076	0.0063	0.0086	0.0082
0.983	2.0	0.0084	0.0149	0.0104	0.0072	0.0251	0.0083	0.0064	0.0118	0.0073
0.648	2.5	0.0092	0.0198	0.0096	0.0078	0.0472	0.0096	0.0066	0.0157	0.0057
0.442	3.0	0.0103	0.0225	0.0085	0.0083	0.0802	0.0110	0.0069	0.0208	0.0052

Table 3: False Alarm rates of the SWV-R, SC-R and WV-R control charts for unknown parameters

Sample size (n)

			5			7			10		
Di	stribution	α3	SWV-R	SC-R	WV-R	SWV-R	SC-R	WV-R	SWV-R	SC-R	WV-R
nc	rmal	0.0	0.0046	0.0045	0.0045	0.0044	0.0045	0.0044	0.0043	0.0041	0.0043
	3.6286	0.0	0.0029	0.0011	0.0029	0.0026	0.0020	0.0026	0.0026	0.0026	0.0026
	2.2266	0.5	0.0029	0.0029	0.0037	0.0030	0.0044	0.0036	0.0032	0.0059	0.0038
β	1.5688	1.0	0.0040	0.0054	0.0057	0.0039	0.0076	0.0055	0.0039	0.0092	0.0055
	1.2123	1.5	0.0047	0.0059	0.0075	0.0045	0.0083	0.0071	0.0043	0.0108	0.0070
	0.9987	2.0	0.0052	0.0039	0.0089	0.0049	0.0070	0.0085	0.0046	0.0092	0.0081
	0.8598	2.5	0.0056	0.0050	0.0100	0.0052	0.0040	0.0094	0.0048	0.0058	0.0090
	0.7637	3.0	0.0058	0.0043	0.0108	0.0054	0.0039	0.0102	0.0049	0.0041	0.0096
ga	mma										
-	38000	0.0	0.0045	0.0014	0.0046	0.0044	0.0019	0.0044	0.0042	0.0024	0.0043
	15.4	0.5	0.0046	0.0023	0.0054	0.0045	0.0033	0.0052	0.0047	0.0045	0.0054
	3.913	1.0	0.0053	0.0045	0.0070	0.0052	0.0061	0.0070	0.0051	0.0071	0.0070
α	1.788	1.5	0.0054	0.0052	0.0082	0.0053	0.0070	0.0080	0.0050	0.0089	0.0077
	0.983	2.0	0.0051	0.0036	0.0087	0.0048	0.0063	0.0081	0.0045	0.0088	0.0081
	0.648	2.5	0.0046	0.0028	0.0089	0.0043	0.0040	0.0086	0.0040	0.0067	0.0082
	0.442	3.0	0.0043	0.0036	0.0094	0.0039	0.0052	0.0133	0.0036	0.0046	0.0084

Bold values are significant

Table 4: Probabilities of out-of-control of variant dispersion control charts, Weibull shape parameter $\beta = 1, n = 5$

n = 5									
Distribution/δ	SWV-R	Exact R-chart	SC-R	WV-R					
Weibull, $\beta = 1$ (dalta)									
1.1	0.0097	0.0039	0.0061	0.0159					
1.2	0.0160	0.0062	0.0095	0.0250					
1.3	0.0244	0.0095	0.0143	0.0363					
1.4	0.0349	0.0143	0.0212	0.0504					
1.5	0.0477	0.0210	0.0303	0.0676					
2.0	0.1385	0.0745	0.0977	0.1770					
2.5	0.2542	0.1593	0.1962	0.3071					
3.0	0.3727	0.2571	0.3025	0.4292					
3.5	0.4796	0.3574	0.4078	0.5382					
4.0	0.5717	0.4496	0.5008	0.6265					
4.5	0.6474	0.5308	0.5802	0.6973					

Table 5: Probabilities of out-of-control of variant dispersion control charts, Weibull shape parameter $\beta = 1$, n = 7

$\Pi = 7$								
Distribution/δ	SWV-R	Exact R-chart	SC-R	WV-R				
Weibull, $\beta = 1$ (dalta)								
1.1	0.0095	0.0038	0.0080	0.0156				
1.2	0.0162	0.0062	0.0109	0.0255				
1.3	0.0254	0.0099	0.0160	0.0385				
1.4	0.0375	0.0154	0.0234	0.0546				
1.5	0.0518	0.0226	0.0329	0.0739				
2.0	0.1615	0.0880	0.1149	0.2070				
2.5	0.3040	0.1925	0.2354	0.3647				
3	0.4473	0.3160	0.3689	0.5121				
3.5	0.5716	0.4373	0.4929	0.6333				
4.0	0.6722	0.5445	0.5989	0.7262				
4.5	0.7504	0.6362	0.6858	0.7964				

Table 6: Probabilities of out- of- control of variant dispersion control charts, Weibull shape parameter $\beta = 1, n = 10$

n = 10

Distribution/8	SWV-R	Exact R-chart	SC-R	WV-R
Weibull, $\beta = 1$ (dalta)			
1.1	0.0094	0.0037	0.0088	0.0155
1.2	0.0165	0.0063	0.0114	0.0262
1.3	0.0269	0.0106	0.0168	0.0407
1.4	0.0404	0.0170	0.0251	0.0596
1.5	0.0572	0.0256	0.0362	0.0817
2.0	0.1883	0.1056	0.1344	0.2412
2.5	0.3622	0.2373	0.2842	0.4315
3	0.5296	0.3886	0.4441	0.6001
3.5	0.5296	0.5333	0.5883	0.7302
4.0	0.6691	0.6525	0.7025	0.8204
4.5	0.7715	0.7471	0.7890	0.8819

skewnesses and sample sizes, when the distributions are Weibull and gamma. Table 4-6 show that the probabilities of out-of-control detections of the proposed SWV-R and SC-R charts are close to those of the exact R chart than the WV-R control chart. In general, the proposed SWV-R control chart provides good performances in term of false alarm rate and probability of out-of-control detection for all levels of skewnesses, sample sizes and magnitudes of shifts.

CONCLUSION

In this study, we have proposed the SWV-R control chart for skewed populations. This proposed chart based on the scaled weighted variance method suggested by Castagliola (2000). The proposed SWV-R control chart reduces to the Shewhart R control chart when the underlying population has a normal distribution. Our simulation study on the false alarm rate indicates that the SWV-R control chart provides lower false alarm rates than those of SC-R and WV-R control charts for all levels of skewnesses and sample sizes. The proposed SWV-R control chart offers considerable improvement over the SC-R and WV-R control charts when it is desirable for the false alarm rate to be closed to the conventional 0.0027. In the case of the probability of out-of-control detections, the simulation results show that the said probabilities of the proposed SWV-R control chart are closer to the chart constructed by exact R chart than the WV-R control charts. The findings are based on the SWV-R method instead of relying on the SC-R and WV-R. Hence, the SWV-R chart can act as a favorable substitute to the existing SC-R and WV-R control charts in the evaluation of the speed of a chart to detect shifts in process dispersion, when the underlying distribution is skewed.

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