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Research Article Stochastic Modelling of Insulin Sensitivity for out of Hospital Cardiac Arrest Patients Treated with Hypothermia

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Abstract

Hypothermia is often used to treat out of hospital cardiac arrest (OHCA) patients who often simultaneously receive insulin for stress induced hyperglycaemia. Variations in response to insulin reflect dynamic changes in insulin sensitivity (SI), defined by the overall metabolic response to stress and therapy. Thus, tracking and forecasting this parameter is important to provide safe glycaemic control in highly dynamic patients. This study examines stochastic forecasting models of model-based SI variability in OHCA patients to assess the resulting potential impact of this therapy on glycaemic control quality and safety. A retrospective analysis of clinically validated model-based SI profiles identified using data from 240 post-cardiac arrest patients (9988 h) treated with hypothermia, shortly after admission in the Intensive Care Unit (ICU). Data were divided into three periods: (1) Cool ($T \ge 35^{\circ}$ C), (2) Idle period of 2 h as hypothermia was removed and (3) Warm ($T \ge 37^{\circ}$ C). The stochastic model captured 60.7 and 90.2% of SI predictions within the (25-75th) and (5-95th) probability forecast intervals during cool period. Equally, it is also recorded 62.8 and 92.1% of SI predictions respectively during the warm period. Maintaining the kernel density variance estimator to c = 1.0 yielded 60.7 and 90.2% for the cool period. Similarly, adjusting a variance estimator of c = 2.0 yields 60.4 and 90.1% for the warm period. A cohort-specific stochastic model of SI provided a conservative forecast for the inter-quartile range and was relatively exact for the 90% range. Adjusting the variance estimator provides a more accurate, cohort-specific stochastic model of SI dynamics for the 90% range. These latter results show clearly different levels and distribution of forecasted S₁ variability between the cold and warm periods.

Key words: OHCA, stochastic method, ICING model, cohort-specific model

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Data Availability: All relevant data are within the paper and its supporting information files.

INTRODUCTION

Hyperglycaemia is prevalent in critical care¹⁻⁴ and increases the risks of further complications and mortality^{1,4,5}. Glycaemic control has shown benefits in reducing mortality^{4,6,7}. However, due in parts to excessive metabolic variability⁸, other studies have found it difficult to reproduce these results⁹⁻¹¹. Out-of-hospital cardiac arrest (OHCA) patients have low survival rates and often experience hyperglycaemia^{12,13}. However, cardiac patients are one group who have more consistently shown benefit from Glycaemic Control (GC), but can be highly insulin resistant and variable, particularly on the first day of stay¹⁴.

Hypothermia is often used to treat OHCA patients¹³. In general, it leads to a lowering of metabolic rate that induces changes in energy metabolism. However, its impact on metabolism and insulin resistance in critical illness is unknown, although one of the adverse events associated with hypothermic therapy is a decrease in insulin sensitivity and insulin secretion¹⁵. However, this decrease may not be notable in a cohort that is already highly resistant and variable¹⁴. Hence, understanding metabolic evolution and variability would enable safer, more accurate GC in this cohort.

Model-based glycaemic control methods using both insulin and/or nutrition modulation have been employed successfully in the control of hyperglycaemia. These methods allow the derivation of patient metabolic state, S_I in this case, by using serial blood glucose (BG) measurements and records of nutrition and insulin administration¹⁶. Once the current S_I has been identified, prediction of future S_I would allow predictions of outcome BG concentration for an intended clinical intervention¹⁷.

Variations in the S_I parameter reflect the metabolic response to stress² and drug therapy¹⁸. Thus, tracking and forecasting this parameter is important to provide safe glycaemic control in the highly dynamic out-of-hospital cardiac arrest (OHCA) patients, treated with hypothermia. Since stochastic modeling has shown its ability to quantify the probability of a future S_I^{17} , the resulting distribution of BG concentrations that would result from a given intervention can be determined^{17,19,20,21}. This information can be used to guide both insulin and/or nutrition interventions, which is the key to avoid unintended hypoglycaemia, improve overall glycaemic control and identify periods of potential high glucose variability that may be indicative of unusual clinical events or cohorts²².

This study presents the adaptation of a stochastic model for S_1 prediction from adult critical care to the unique clinical and physiological case of OHCA patients, treated with

hypothermia. The stochastic model presented is constructed by the distribution of insulin sensitivity variation using a 2-D kernel density method. Clinically validated, model-based insulin sensitivity (S₁)^{20,23} is used to provide more accurate measure of patient metabolic state and its stochastic model during cool and warm periods. Modifications to the initial kernel density estimation model are made to explore and optimize the relationship between the model and the underlying dataset, particularly for the clinically relevant 90% range that can be used to quantify and minimize the risks of hyper or hypo-glycaemia as an integral part of a control protocol^{20,21}.

MATERIALS AND METHODS

Patients and data: A retrospective analysis of glycaemic control data from 240 OHCA patients (9988 h) treated with hypothermia, shortly after admission to the Intensive Care Unit (ICU) at Christchurch Hospital, New Zealand, Erasme Hospital, Belgium and Lausanne Hospital, Switzerland. Patients from Christchurch Hospital (20) were on the SPRINT glycaemic control protocol⁷, whereas the remaining 160 patients from Erasme (82) and Lausanne (78) Hospitals were on unpublished local glycaemic protocols. Blood Glucose (BG) and temperature readings were taken 1-2 h along with all insulin and nutrition data. Data were divided into three periods: (1) Cool (T \leq 35°C) (2) Transition period of 2 h as hypothermia was removed and (3) Warm after (T \ge 37°C). A maximum of 24 h and a minimum of 15 h for each period (1 and 3) were considered, ensuring a balance of contiguous data between periods. Overall demographics are shown in Table 1.

Therapeutic hypothermia: Therapeutic Hypothermia (TH) was applied following a standardized written protocol. All patients were treated with mild TH to 33 ± 1 °C for up to 24 h, irrespective of age, initial arrest rhythm and other physiological conditions. The TH was started immediately after admission and was induced with ice-cold packs and/or intravenous cold fluids. Body temperature was maintained

Table 1: Demographic	data for al	l patients
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	Value	Value					
N/ · · I I							
Variable	Cool	Warm					
Total patients (n)	240	240					
Total treatment (h)	4987	5001					
Blood glucose (mmol L ⁻¹) (IQR)	7.40 [6.20-9.70]	6.56 [5.61-7.78]					
Insulin rate (U h ⁻¹) (IQR)	3.37 [1.33-8.00]	3.51 [1.60-7.00]					
Glucose rate (g h^{-1}) (IQR)	2.69 [1.04-5.26]	5.41 [2.71-8.11]					

Data are presented as median (interquartile range) where appropriate

at hypothermia using a surface cooling device with a computerized adjustment of patient temperature target. During this time, some short-acting drugs, such as midazolam (0.1 mg kg⁻¹ h⁻¹), fentanyl (1.5 μ g kg⁻¹ h⁻¹) and vecuronium (0.1 mg kg⁻¹ boluses) were used to administer sedation, analgesia and control shivering. Rewarming was achieved passively and sedation-analgesia was typically stopped when patient temperature was greater than 35 °C.

Metabolic system model: Model-based insulin sensitivity (S₁) in this study is a patient-specific parameter describing the whole body effect of insulin. The analysis of patient-specific insulin sensitivity employs the ICING model²⁴ as a glucose-insulin system model developed and clinically validated in critical care glycaemic control and insulin sensitivity studies^{20,21,23,24}. It is defined as:

$$G = -p_{G}.G(t) - S_{t}.G(t).\frac{Q(t)}{1 + \alpha_{G}Q(t)} + \frac{P(t) + EGP_{b} - CNS}{V_{G}}$$
(1)

$$i = -n_{K}I(t) - \frac{n_{L}I(t)}{1 + \alpha_{I}I(t)} - n_{I}(I(t) - Q(t)) + \frac{u_{ex}(t)}{V_{I}} + (1 - X_{L})\frac{u_{en}(G)}{V_{I}}$$
(2)

$$\dot{Q} = n_1 (I(t) - Q(t)) - n_C \frac{Q(t)}{1 + \alpha_G Q(t)}$$
 (3)

$$\dot{P}_1 = d_1 P_1 + D(t)$$
 (4)

Table 2: Parameter values and descriptions for the ICING model

$$P_2 = -\min(d_2P_2, P_{max}) + d_1P_1$$
 (5)

$$P(t) = \min\left(d_2 P_2, P_{max}\right) + PN(t)$$
(6)

$$\mathbf{u}_{en}(\mathbf{G}) = \min \max \left(\mathbf{u}_{\min}, \mathbf{k}_1 \mathbf{G}(t) + \mathbf{k}_2 \right), \mathbf{u}_{\max}$$
(7)

A summary of parameter values and descriptions and exogenous input variables for the model are listed in Table 2 and 3, respectively.

Insulin sensitivity S₁ is identified hourly from patient data, producing a step-wise hourly varying profile²⁵. This profile effectively describes patient-specific metabolic behavior under time-varying physiologic conditions. The validity and independence of this patient-specific parameter have been validated using data from independent, clinically matched cohorts²³ and in gold-standard insulin sensitivity tests²⁶.

Stochastic model: A 2-D kernel density estimation method is used to construct the stochastic model that describes the hourly transition of S₁. The kernel density method combines probability distribution functions for each point of data to generate an overall density function for the dataset. This method has the advantage of producing a smooth, physiologically likely, continuous function across the parameter range to provide continuity when interpolating S₁ forecasts to account for each particular patient state. It also

Parameter	Value	Description
p _G	0.006 (min ⁻¹)	Non-insulin mediated glucose removal
EGP	1.16 (mmol min ⁻¹)	Endogenous glucose production rate
CNS	0.3 (mmol min ⁻¹)	Central nervous system glucose uptake
V _G	13.3 (L)	Plasma glucose distribution volume
V	4.0 (L)	Plasma and interstitial insulin distribution volume
α _G	0.0154 (L mU ⁻¹)	Insulin binding saturation parameter
αι	0.0017 (L mU ⁻¹)	Hepatic insulin clearance saturation parameter
n _l	0.006 (min ⁻¹)	Trans-endothelial diffusion rate
n _c	0.006 (min ⁻¹)	Interstitial insulin degradation rate
n _K	0.0542 (min ⁻¹)	Renal insulin clearance rate
n _L	0.1578 (min ⁻¹)	Hepatic insulin clearance rate
XL	0.67	Fractional first-pass hepatic insulin extraction
d ₁	0.0347 (min ⁻¹)	Glucose transport rate from stomach to gut
d ₂	0.0069 (min ⁻¹)	Glucose transport rate from gut to plasma
P _{max}	6.11 (mmol min ⁻¹)	Maximum glucose flux from gut to plasma
U _{min}	16.7 (mU min ⁻¹)	Minimum pancreatic secretion rate
U _{max}	266.7 (mU min ⁻¹)	Maximum pancreatic secretion rate
k ₁	ND 14.9 (mU L mmol ⁻¹ min ⁻¹)	Pancreatic insulin secretion glucose-sensitivity
	T2DM 4.9	
	T1DM 0.0	
k ₂	ND-49.9 (mU min ⁻¹)	Pancreatic insulin secretion offset
	T2DM -27.4	
	T1DM 16.7	

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Variable	Description
PN (t) (mmol min ⁻¹)	Intravenous glucose input rate (parenteral nutrition)
D (t) (mmol min ⁻¹)	Oral glucose input rate (enteral nutrition)
u _{ex} (t) (mU min ⁻¹)	Intravenous insulin input rate

Table 4: Descriptions of 6 h blocks for data analysis

Table 3. Exogenous model input variables

	6 h blocks		
Day	Block	Hours range (h)	Period
1	1	0-6	Cool
	2	6-12	Cool
	3	12-18	Cool
	4	18-24	Cool
2	5	24-30	Warm
	6	30-36	Warm
	7	36-42	Warm
	8	42-48	Warm

automatically accounts for any possible multimodality where the density of data may show several distinct peaks corresponding to patterns of changes in S_I. The overall result is a bivariate probability density function for the potential parameter values. The goal of this statistical model is to quantify the range of S_I 1 h ahead in time (S_{I, n+1}) based on available data (S_{I,n}, S_{I,n-1}, S_{I,n-2}, ..., S_{I,0}) to guide real-time clinical control^{22,21}. Thus, it is potentially important that the model is also as cohort-specific as possible for greatest accuracy and to minimize over-conservative forecasts²².

A 2-D kernel density method is chosen because the distribution of $S_{l,n+1}$ varies with $S_{l,n}$ and cannot be simply described with a single standard statistical distribution. Thus, the variations in S_l can be treated as a Markov process. A Markov process has the property that the conditional probability density function of future states of the process, given the current state, depends only upon the current state. Therefore, using the Markov property of the stochastic behaviour of S_l , the conditional probability density of $S_{l,n+1}$ taking on a value y can be calculated by knowing $S_{l,n} = x$. Model equations and derivation were defined¹⁷.

In this study, the model was cross validated by splitting the 240 of cool and warm patient cohorts into five groups, each containing 48 patients per period following the method of Lin *et al.*¹⁷. For each group, the model created using the remaining 192 patients of the cohort representing approximately 3990 h of data per period. Out-of-sample S_I predictions were generated for the 48 patients of unused group and compared to the actual fitted S_I from these 48 patients to assess model accuracy and cohort generality. Based on results from in-sample tests, where the stochastic model is generated from the entire retrospective dataset and tested on the same data and out-of-sample tests, where different subsets of data are used for model generation and testing, the kernel density estimator was modified by multiplying the variance estimators by a constant c (i.e., $c\sigma_x$ and $c\sigma_y$) to explore the model bias-variance trade-off for this data between cool and warm periods as well as 6 and 12 h blocks. This adjustment to the variance estimator effectively adjusts the kernel bandwidth and the degree of smoothing over the data, which is used here to assess variability differences between periods or cohorts.

Analyses and metrics: Current S₁ (n) during the cool (T \leq 35 °C) and warm (T \geq 37 °C) periods were identified hourly using the ICING model²⁴ for each patient. Using current S₁ (n) data, the predicted S₁ (n+1) data can be generalized by shifting 1 h forward from the current S₁ (n) data. These S₁ vectors will be paired and used to create the stochastic S₁ model of OHCA cohort. Stochastic S₁ models were analyzed during both cool and warm periods as follows:

- Overall cohort patient
- Analysis of patients in 6 h block

Overall cohort analysis assessed the stochastic model behavior of insulin sensitivity during both cool and warm periods, which includes percentage of S_1 within prediction interval and analysis of modifying kernel density estimation. The S_1 is also analyzed using 6 h blocks as described in Table 4, to capture S_1 variability and forecasting over time with different resolution. The analysis includes percentage of S_1 within predicted interval and analysis of modifying the kernel density estimation.

RESULTS AND DISCUSSION

Distribution of hourly variation in S₁: Figure 1 presents the distribution of hourly variation in S₁ and the stochastic model percentile probability bands at c = 1.0 for the 240 OHCA patients during cool (4987 h) and warm (5001 h) periods. Approximately 85% of the values during cool period and 70% during warm period are below $1.0 \times 10-3$ L mU⁻¹ min⁻¹. The results show that the hourly variation of S₁ is wider during the cool period and S₁ rises from cool to warm.

Table 5 shows the in-sample results of stochastic model prediction widths for both cool ($n_{cool} = 4622$ predictions) and warm ($n_{warm} = 4832$ predictions) periods. The number of predictions is less than the total hours of S₁ as the patient data records are not always perfectly divisible by 1 h and because predictions can only be computed after the second hour of patient data.

For the cool period, S₁ predictions (60.7%) were within the (25-75th) probability intervals and within the (5-95th) probability interval. Thus, the proportion of S₁ values that fell within the forecast (25-75th) probability intervals were measurably higher than the expected 50%, but very much closer for the 90% intervals. Similar results could be observed for the warm period, but were measured higher than cool period for both 50 and 90% intervals.

The percentage S_1 within prediction interval results of Lin *et al.*²⁴ of 54.0% is within the (25-75th) probability bound shows that normal adult ICU patients with normal body temperature produce far closer to the ideal 50%. However, the prediction interval percentage results of 60.7 and 62.8% at cool and warm respectively on the same probability bound for the OHCA patients, treated with hypothermia as shown in Table 5 appear to be unique and significant. These results are similar to neonate's results of Le Compte *et al.*¹⁹, which record 62.6% at the same interval.

Cross-validation comparison studies: Table 6 shows the results of the cross validation comparison study for 240 patient's cohort during the cool and warm periods,

Table 6: Cross-validation comparison study for 240 patient cohort

respectively. Generally, these results are consistent between groups, suggesting that the overall model contains sufficient data to account for the range of dynamics observed in this cohort.

Probability-bound determination using local variance estimator: Table 7 shows the effect of modifying the kernel density estimation for several values of c, ranging from 0.1-3.0. For this cohort, the increase of c>1.0, yield better coverage widths for the expected proportions. Thus, for cool period, the value c = 1.0 enables the best 90% interval coverage, but is conservative for the inter-quartile range. Similarly, c = 2.0 is best for the warm period.

Table 5: Sample results for in-sample stochastic model prediction widths at c = 1

		Value							
Variable	Prediction width/range	 Cool (%)	Warm (%)						
Percentage of S _I within	25-75th	60.7	62.8						
prediction interval	5-95th	90.2	92.1						
Data are presented as cohort median ($n_{cool} = 4622$ predictions and $n_{warm} = 4832$									

predictions)

		Percentage of S _I within in	terval during cool period	Percentage of S_1 within interval during warm peric		
Group	Groups used to create the model	 25-75th	5-95th	 25-75th	5-95th	
1	[-, 2, 3, 4, 5]	60.1	90.5	64.1	92.3	
2	[1, -, 3, 4, 5]	61.2	90.0	64.0	92.6	
3	[1, 2, -, 4, 5]	61.9	90.8	61.7	91.8	
4	[1, 2, 3, -, 5]	61.9	90.8	63.1	92.7	
5	[1, 2, 3, 4, -]	62.6	90.7	63.8	92.8	
Overall	[1, 2, 3, 4, 5]	60.7	90.2	62.8	92.1	



Fig. 1(a-b): Probability interval and distribution of hourly variation in S₁ for OHCA patients, treated with hypothermia (a) During cool and (b) Warm (right) periods



Fig. 2(a-b): Probability-bound determination for raw S₁ data and corresponding BG forecasted values based on an equal-tailed 0.90 probability interval of SI between 5th -95th at local variance estimator, both cool (c = 1.0) and warm (c = 2.0) period. The solid lines represent the 5, 25, 50, 75 and 95% probability bounds

С	Percentage of S ₁ within proba	pility bound during cool period	Percentage of S_1 within probability bound during warm period				
	 25-75th	5-95th	 25-75th	 5-95th			
0.1	49.3	87.8	45.5	83.7			
0.2	50.2	88.4	48.2	84.4			
0.3	51.1	88.5	50.4	85.2			
0.5	53.5	88.6	52.2	85.9			
1.0	60.7	90.2	55.8	87.1			
1.5	66.2	91.3	58.0	88.9			
2.0	69.6	92.1	60.4	90.1			
2.5	72.2	92.7	62.8	91.2			
3.0	74.0	93.4	64.9	91.5			
Ideal	50%	90%	50%	90%			

Table 7: Comparison of probability bounds for modifications of kernel density estimator ($\sigma_x = c\sigma_x$ and $\sigma_y = c\sigma_y$) during both cool and warm periods

The difference of optimal c values between cool and warm suggests that the variation and stochastic modeling for both periods are different, leading to potentially different control requirements to ensure safe glycaemic control in these highly dynamic patients and states. In particular, c = 1.0 is currently used in the STAR protocol^{27,28} in medical ICU patients. Hence, no change is required for the cool period. However, during the warm period, the stochastic model requires a wider and smoother probability distribution coverage at c = 2.0, which would necessitate a control change of this known change of state. These results are further reflected in Fig. 2.

The kernel density estimator method employed in this stochastic model provides a layer of safety as wider probability bounds would be more likely to capture dynamics and any changes not observed in the cohort. As the (5-95th) band is what has been used for control previously, these cohorts show that they are closer to ideal 90%. However, wider coverage bands may also have impact on glycaemic control performance. As the wider probability band might be useful to avoid potential hypoglycaemia, it may also force a controller to maintain a mildly hyperglycaemic state. Importantly, different level of smoothing are required for the cool and warm periods, indicating that cool patients are less multi-modal (more smoothed) and warm OHCA patients are more multi-modal (less smoothed), in comparison to broad ICU cohorts. Perhaps this scenario is caused by the suppression of metabolic activities during cool period.

Stochastic analysis by 6 h block: The analysis of stochastic model based on 6 h time block have used the same methodology as the overall cohort patient, except that the patient cohort data are separated in block hours. Table 8 presents the in-sample results of stochastic model prediction widths based on 6 hour block analysis at c = 1.0.

The results show that S_1 prediction interval coverage meets the prediction expectation despite exceeding its ideal

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Table 8: Stochastic model prediction widths based on 6 h block analysis

		Cool				Warm	Warm				
		Block 1	Block 2	Block 3	Block 4	Block 5	Block 6	Block 7	Block 8		
Variable	Prediction width/range	(0-6 h)	(6-12 h)	(12-18 h)	(18-24 h)	(24-30 h)	(30-36 h)	(36-42 h)	(42-48 h)		
SI within	25-75th	56.1	60.2	58.9	61.5	61.9	62.1	62.6	62.6		
prediction	5-95th	90.1	91.9	91.0	91.3	91.2	92.2	92.3	92.7		

Table 9: Comparison of probability bounds for modifications of kernel density estimator ($\sigma'_x = c\sigma_x$ and $\sigma'_y = c\sigma_y$) based on 6 h block analysis

	Cool period								Warm period							
	Percentage of S ₁ within probability bounds at block 1 (0-6 h)		Percenta within pr bounds a block 2 (6	ge of SI robability at 5-12 h)	Percenta within pr bounds a block 3 (ge of SI robability at 12-18 h)	Percenta within pr bounds a block 4 (1	ge of SI obability It I8-24 h)	Percenta within pr bounds a block 5 (2	ge of SI robability at 24–36 h)	Percenta within pr bounds a block 6 (3	ge of SI obability it 36-48 h)	Percenta within pr bounds a block 7 (2	ge of SI robability at 36-48 h)	Percentage of SI within probability bounds at block 8 (36-48 h)	
С	25 -75th	5-95th	25-75th	5-95th	25-75th	5-95th	25-75th	5-95th	25-75th	5-95th	25-75th	5-95th	25-75th	5-95th	25-75th	5-95th
0.1	51.6	88.9	52.7	85.3	53.5	85.8	51.7	86.0	44.9	84.2	45.0	85.3	45.5	85.5	46.2	85.6
0.2	51.4	89.3	53.5	86.2	54.3	86.6	52.1	86.2	45.3	84.5	45.4	85.8	46.2	85.8	46.7	85.9
0.3	51.7	89.5	54.0	86.9	54.7	86.9	53.5	86.6	46.3	85.8	46.7	86.6	47.4	86.6	48.0	86.7
0.5	53.1	89.6	55.6	88.8	55.1	87.8	54.3	87.4	48.4	86.5	48.6	87.0	49.2	87.1	49.3	87.3
1.0	58.7	90.4	57.2	89.3	56.9	89.0	55.5	87.3	51.9	87.2	52.1	88.2	52.6	88.3	52.6	87.7
1.5	62.7	91.0	58.9	90.7	58.0	90.6	57.2	88.2	54.9	87.8	55.2	88.6	55.5	88.8	55.7	88.3
2.0	65.6	91.5	61.0	93.5	60.2	92.1	59.2	90.7	55.6	90.3	56.1	89.3	56.2	89.4	57.2	88.9
2.5	67.3	92.3	63.4	94.2	62.7	92.5	60.8	92.4	56.2	91.9	55.8	90.9	57.3	90.1	58.5	89.5
3.0	68.0	92.8	65.7	94.9	63.9	92.6	62.1	92.9	56.7	92.8	57.5	91.2	57.9	91.3	59.0	90.7
Ideal(%) 50	90	50	90	50	90	50	90	50	90	50	90	50	90	50	90

values of 50 and 90% for both (25-75th) and (5-95th) interval width respectively. It is also observed that the proportion of fitted S_1 values that fell within the (25-75th) and (5-95th) probability intervals for 6 h block analysis were measured higher than the expected 50 and 90%.

Table 9 shows the effect of modifying the kernel density estimation for several values of c, ranging from 0.1-3.0 based on 6 h block analysis. For this cohort, the results has shown that for block 1, c = 1.0 and followed by block 2-3 (c = 1.5), block 4-5 (c = 2.0), block 6-7 (c = 2.5) and block 8 (c = 3.0). The trend shows that the value of estimator, c is increased as S_I increases from cool to warm and match with overall cohort stochastic model analysis as shown in the Table 7. This will lead to another idea of separating stochastic model for each time block to implement stochastic control. The difference of optimal c values between 6 h blocks suggests that the variation and stochastic modeling for each time block is different, leading to different control requirements to ensure safe glycaemic control in the highly dynamic conditions.

Overall, this stochastic method and analysis in this study provides predictions based on a cohort dataset. The prediction bounds for more dynamic patients are difficult to decide since the SI level and variability distribution for this cohort is unique²⁹ and tracked its evolution over time, particularly during cool period. This observation is far differing than for the less dynamic patients who are typically more conservative.

CONCLUSION

Thus, in this study the probability bounds are optimized in a cohort sense, but not necessarily applicable on a per-patient basis.

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REFERENCES

- Capes, S.E., D. Hunt, K. Malmberg and H.C. Gerstein, 2000. Stress hyperglycaemia and increased risk of death after myocardial infarction in patients with and without diabetes: A systematic overview. Lancet, 355: 773-778.
- McCowen, K.C., A. Malhotra and B.R. Bistrian, 2001. Stress-induced hyperglycemia. Critical Care Clin., 17:107-124.
- Mizock, B.A., 2001. Alterations in fuel metabolism in critical illness: Hyperglycaemia. Best Pract. Res. Clin. Endocrinol. Metab., 15: 533-551.

- 4. Van den Berghe, G., P. Wouters, F. Weekers, C. Verwaest and F. Bruyninckx *et al.*, 2001. Intensive insulin therapy in critically ill patients. N. Engl. J. Med., 345: 1359-1367.
- 5. Krinsley, J.S., 2003. Association between hyperglycemia and increased hospital mortality in a heterogeneous population of critically ill patients. Mayo. Clin. Proc., 78: 1471-1478.
- 6. Krinsley, J.S., 2004. Effect of an intensive glucose management protocol on the mortality of critically ill adult patients. Mayo Clin. Proc., 79: 992-1000.
- Chase, J., G. Shaw, A. Le Compte, T. Lonergan and M. Willacy *et al.*, 2008. Implementation and evaluation of the SPRINT protocol for tight glycaemic control in critically ill patients: A clinical practice change. Critical Care, Vol. 12. 10.1186/cc6868.
- Chase, J.G., A.J. Le Compte, F. Suhaimi, G.M. Shaw and A. Lynn *et al.*, 2011. Tight glycemic control in critical care-the leading role of insulin sensitivity and patient variability: A review and model-based analysis. Comput. Methods Programs Biomed., 102: 156-171.
- Brunkhorst, F.M., C. Engel, F. Bloos, A. Meier-Hellmann and M. Ragaller *et al.*, 2008. Intensive insulin therapy and pentastarch resuscitation in severe sepsis. N. Engl. J. Med., 358: 125-139.
- Finfer, S., D. R. Chittock, S.Y.S. Su, D. Blair and D. Foster *et al.*, 2009. Intensive versus conventional glucose control in critically ill patients. N. Engl. J. Med., 360: 1283-1297.
- 11. Preiser, J.C., P. Devos, S. Ruiz-Santana, C. Melot and D. Annane *et al.*, 2009. A prospective randomised multi-centre controlled trial on tight glucose control by intensive insulin therapy in adult intensive care units: The Glucontrol study. Intensive Care Med., 35: 1738-1748.
- 12. Taylor, N.A.S., R.F. Griffiths and J.D. Cotter, 1994. Epidemiology of hypothermia: Fatalities and hospitalisations in New Zealand. Aust. N. Z. J. Med., 24: 705-710.
- 13. Neumar, R.W., J.P. Nolan, C. Adrie, M. Aibiki and R.A. Berg et al., 2008. Post-cardiac arrest syndrome: Epidemiology, pathophysiology, treatment and prognostication a consensus statement from the international liaison committee on resuscitation (American heart association, Australian and New Zealand council on resuscitation, European resuscitation council, Heart and stroke foundation of Canada, InterAmerican heart foundation, resuscitation council of Asia and the resuscitation council of Southern Africa): The American heart association emergency cardiovascular care committee: The council on cardiovascular surgery and Anesthesia: The council on cardiopulmonary, perioperative and critical care: The council on clinical cardiology: And the stroke council. Circulation, 118:2452-2483.

- Pretty, C.G., A.J. Le Compte, J.G. Chase, G.M. Shaw, J.C. Preiser, S. Penning and T. Desaive, 2012. Variability of insulin sensitivity during the first 4 days of critical illness: implications for tight glycemic control. Ann. Intensive Care, Vol. 2. 10.1186/2110-5820-2-17.
- Hayashi, N., 2009. Management of pitfalls for the successful clinical use of hypothermia treatment. J. Neurotrauma, 26: 445-453.
- Chase, J.G., G.M. Shaw, T. Lotz, A. LeCompte and J. Wong *et al.*, 2007. Model-based insulin and nutrition administration for tight glycaemic control in critical care. Curr. Drug Delivery, 4: 283-296.
- Lin, J., D. Lee, J.G. Chase, G.M. Shaw and A. Le Compte *et al.*, 2008. Stochastic modelling of insulin sensitivity and adaptive glycemic control for critical care. Comput. Methods Programs Biomed., 89: 141-152.
- Pretty, C., J.G. Chase, J. Lin, G.M. Shaw, A. Le Compte, N. Razak and J.D. Parente, 2011. Impact of glucocorticoids on insulin resistance in the critically ill. Comput. Methods Programs Biomed., 102: 172-180.
- Le Compte, A.J., D.S. Lee, J.G. Chase, J. Lin, A. Lynn and G.M. Shaw, 2010. Blood glucose prediction using stochastic modeling in neonatal intensive care. IEEE Trans. Biomed. Eng., 57: 509-518.
- Evans, A., G.M. Shaw, A. Le Compte, C.S. Tan and L. Ward *et al.*, 2011. Pilot proof of concept clinical trials of Stochastic Targeted (STAR) glycemic control. Anna. Intensive Care, Vol. 1. 10.1186/2110-5820-1-38.
- 21. Fisk, L.M., A.J. Le Compte, G.M. Shaw, S. Penning, T. Desaive and J.G. Chase, 2012. STAR development and protocol comparison. IEEE Trans. Biomed. Eng., 59: 3357-3364.
- Thomas, F., C.G. Pretty, L. Fisk, G.M. Shaw, J.G. Chase and T. Desaive, 2014. Reducing the impact of insulin sensitivity variability on glycaemic outcomes using separate stochastic models within the STAR glycaemic protocol. Biomed. Eng. Online, Vol. 13. 10.1186/1475-925X-13-43
- 23. Chase, J.G., F. Suhaimi, S. Penning, J.C. Preiser and A.J. Le Compte *et al.*, 2010. Validation of a model-based virtual trials method for tight glycemic control in intensive care. Biomed. Eng., Vol. 9. 10.1186/1475-925X-9-84.
- Lin, J., N.N. Razak, C.G. Pretty, A. Le Compte and P. Docherty *et al.*, 2011. A physiological Intensive Control Insulin-Nutrition-Glucose (ICING) model validated in critically ill patients. Comput. Methods Programs Biomed., 102: 192-205.
- Hann, C.E., J.G. Chase, J. Lin, T. Lotz, C.V. Doran and G.M. Shaw, 2005. Integral-based parameter identification for long-term dynamic verification of a glucose-insulin system model. Comput. Methods Programs Biomed., 77: 259-270.

- 26. McAuley, K.A., J.E. Berkeley, P.D. Docherty, T.F. Lotz and L.A. Te Morenga *et al.*, 2011. The dynamic insulin sensitivity and secretion test-a novel measure of insulin sensitivity. Metabolism, 60: 1748-1756.
- 27. Evans, A., A. Le Compte, C.S. Tan, L. Ward and J. Steel *et al.*, 2012. Stochastic targeted (STAR) glycemic control: Design, safety and performance. J. Diabetes Sci. Technol., 6: 102-115.
- Penning, S., A.J. Le Compte, K.T. Moorhead, T. Desaive and P. Massion *et al.*, 2012. First pilot trial of the STAR-Liege protocol for tight glycemic control in critically ill patients. Comput. Methods Programs Biomed., 108: 844-859.
- Pri, A.S., J.G. Chase, C.G. Pretty, G.M. Shaw and J.C. Preiser *et al.*, 2014. Evolution of insulin sensitivity and its variability in out-Of-Hospital Cardiac Arrest (OHCA) patients treated with hypothermia. Critical Care, Vol. 18. 10.1186/s13054-014-0586-x.